# Homework Set #1Entropy rate, directed information

# 1. Monotonicity of entropy per element.

For a stationary stochastic process  $X_1, X_2, \ldots, X_n$ , show that

$$\frac{H(X_1, X_2, \dots, X_n)}{n} \le \frac{H(X_1, X_2, \dots, X_{n-1})}{n-1}.$$

(b)

(a)

$$\frac{H(X_1, X_2, \dots, X_n)}{n} \ge H(X_n | X_{n-1}, \dots, X_1).$$

# Solution: Monotonicity of entropy per element.

(a) By the chain rule for entropy,

$$\frac{H(X_1, X_2, \dots, X_n)}{n} = \frac{\sum_{i=1}^n H(X_i | X^{i-1})}{n} \\
= \frac{H(X_n | X^{n-1}) + \sum_{i=1}^{n-1} H(X_i | X^{i-1})}{n} \\
= \frac{H(X_n | X^{n-1}) + H(X^{n-1})}{n}.$$
(1)

From stationarity it follows that for all  $1 \le i \le n$ ,

$$H(X_n|X^{n-1}) \le H(X_i|X^{i-1}),$$

which further implies, by averaging both sides, that,

$$H(X_n|X^{n-1}) \leq \frac{\sum_{i=1}^{n-1} H(X_i|X^{i-1})}{n-1} \\ = \frac{H(X^{n-1})}{n-1}.$$
(2)

Combining (1) and (2) yields,

$$\frac{H(X_1, X_2, \dots, X_n)}{n} \leq \frac{1}{n} \left[ \frac{H(X_1, X_2, \dots, X_{n-1})}{n-1} + H(X_1, X_2, \dots, X_{n-1}) \right]$$
$$= \frac{H(X_1, X_2, \dots, X_{n-1})}{n-1}.$$

(b) By stationarity we have for all  $1 \le i \le n$ ,

$$H(X_n|X^{n-1}) \le H(X_i|X^{i-1}),$$

which implies that,

$$H(X_n|X^{n-1}) = \frac{\sum_{i=1}^n H(X_n|X^{n-1})}{n}$$
  
$$\leq \frac{\sum_{i=1}^n H(X_i|X^{i-1})}{n}$$
  
$$= \frac{H(X_1, X_2, \dots, X_n)}{n}.$$

#### 2. Pairwise independence.

Let  $X_1, X_2, \ldots, X_{n-1}$  be i.i.d. random variables taking values in  $\{0, 1\}$ , with  $\Pr\{X_i = 1\} = \frac{1}{2}$ . Let  $X_n = 1$  if  $\sum_{i=1}^{n-1} X_i$  is odd and  $X_n = 0$  otherwise. Let  $n \ge 3$ .

- (a) Show that  $X_i$  and  $X_j$  are independent, for  $i \neq j, i, j \in \{1, 2, ..., n\}$ .
- (b) Find  $H(X_i, X_j)$ , for  $i \neq j$ .
- (c) Find  $H(X_1, X_2, \ldots, X_n)$ . Is this equal to  $nH(X_1)$ ?

#### Solution: Pairwise Independence.

 $X_1, X_2, \ldots, X_{n-1}$  are i.i.d. Bernoulli(1/2) random variables. We will first prove that for any  $k \leq n-1$ , the probability that  $\sum_{i=1}^{k} X_i$  is odd is 1/2. We will prove this by induction. Clearly this is true for k = 1. Assume that it is true for k - 1. Let  $S_k = \sum_{i=1}^{k} X_i$ . Then

$$P(S_k \text{ odd}) = P(S_{k-1} \text{ odd})P(X_k = 0)$$

$$+ P(S_{k-1} \text{ even})P(X_k = 1)$$
(3)

$$= \frac{1}{22} + \frac{1}{22} + \frac{1}{22}$$
(4)

$$= \frac{1}{2}.$$
 (5)

Hence for all  $k \leq n-1$ , the probability that  $S_k$  is odd is equal to the probability that it is even. Hence,

$$P(X_n = 1) = P(X_n = 0) = \frac{1}{2}.$$
(6)

(a) It is clear that when i and j are both less than n,  $X_i$  and  $X_j$  are independent. The only possible problem is when j = n. Taking i = 1 without loss of generality,

$$P(X_1 = 1, X_n = 1) = P(X_1 = 1, \sum_{i=2}^{n-1} X_i \text{ even})$$
 (7)

$$= P(X_1 = 1)P(\sum_{i=2}^{n-1} X_i \text{ even})$$
 (8)

$$= \frac{1}{2} \frac{1}{2} \tag{9}$$

$$= P(X_1 = 1)P(X_n = 1)$$
(10)

and similarly for other possible values of the pair  $X_1, X_n$ . Hence  $X_1$  and  $X_n$  are independent.

(b) Since  $X_i$  and  $X_j$  are independent and uniformly distributed on  $\{0, 1\},\$ 

$$H(X_i, X_j) = H(X_i) + H(X_j) = 1 + 1 = 2$$
 bits. (11)

(c) By the chain rule and the independence of  $X_1, X_2, \ldots, X_{n_1}$ , we have

$$H(X_1, X_2, \dots, X_n) = H(X^{n-1}) + H(X_n | X^{n-1})$$
(12)

$$= \sum_{i=1}^{N-1} H(X_i) + 0 \tag{13}$$

$$= n-1, \tag{14}$$

since  $X_n$  is a function of the previous  $X_i$ 's. The total entropy is not n, which is what would be obtained if the  $X_i$ 's were all independent. This example illustrates that pairwise independence does not imply complete independence.

# 3. Cesáro mean.

Prove that if  $\lim a_n = a$  and  $b_n = \sum_{i=1}^n a_i$ , then  $\lim b_n = a$ .

# Solution: Cesáro mean.

Let  $\varepsilon > 0$ . Since  $a_n \to a$ , there exists a number  $N(\varepsilon)$  such that  $|a_n - a| \leq \varepsilon$  for all  $n \geq N(\varepsilon)$ . Furthermore,

$$b_n - a| = \left|\frac{1}{n}\sum_{i=1}^n (a_i - a)\right|$$

$$\leq \frac{1}{n}\sum_{i=1}^n |(a_i - a)|$$

$$\leq \frac{1}{n}\sum_{i=1}^{N(\varepsilon)} |(a_i - a)| + \frac{n - N(\varepsilon)}{n}\varepsilon$$

$$\leq \frac{1}{n}\sum_{i=1}^{N(\varepsilon)} |(a_i - a)| + \varepsilon$$

for all  $n \ge N(\varepsilon)$ . Since the first term goes to 0 as  $n \to \infty$ , we can make  $|b_n - a| \le 2\varepsilon$  by taking n large enough. Hence,  $b_n \to a$  as  $n \to \infty$ .

#### 4. Stationary processes.

Let  $\ldots, X_{-1}, X_0, X_1, \ldots$  be a stationary (not necessarily Markov) stochastic process. Which of the following statements are true? State true or false. Then either prove or provide a counterexample. Warning: At least one answer is false.

- (a)  $H(X_n|X_0) = H(X_{-n}|X_0)$ .
- (b)  $H(X_n|X_0) \ge H(X_{n-1}|X_0)$ .
- (c)  $H(X_n|X_1^{n-1}, X_{n+1})$  is nonincreasing in n.

#### Solution: Stationary processes.

(a)  $H(X_n|X_0) = H(X_{-n}|X_0).$ 

This statement is true, since

$$H(X_n|X_0) = H(X_n, X_0) - H(X_0)$$
(15)

$$H(X_{-n}|X_0) = H(X_{-n}, X_0) - H(X_0)$$
(16)

and  $H(X_n, X_0) = H(X_{-n}, X_0)$  by stationarity.

(b)  $H(X_n|X_0) \ge H(X_{n-1}|X_0).$ 

Solution: This statement is not true in general, though it is true for first order Markov chains. A simple counterexample is a periodic process with period n. Let  $X_0, X_1, X_2, \ldots, X_{n-1}$  be i.i.d. uniformly distributed binary random variables and let  $X_k = X_{k-n}$ for  $k \ge n$ . In this case,  $H(X_n|X_0) = 0$  and  $H(X_{n-1}|X_0) = 1$ , contradicting the statement  $H(X_n|X_0) \ge H(X_{n-1}|X_0)$ .

(c)  $H(X_n|X_1^{n-1}, X_{n+1})$  is non-increasing in n.

This statement is true, since by stationarity  $H(X_n|X_1^{n-1}, X_{n+1}) = H(X_{n+1}|X_2^n, X_{n+2}) \ge H(X_{n+1}|X_1^n, X_{n+2})$  where the inequality follows from the fact that conditioning reduces entropy.

# 5. Directed Information and causal conditioning

Directed information is denoted as  $I(X^n \to Y^n)$  and is defined as

$$I(X^n \to Y^n) \triangleq \sum_{i=1}^n I(X^i; Y_i | Y^{i-1}).$$

Causal conditioning is denoted as  $p(y^n||x^{n-d})$  and is defined as

$$p(y^{n}||x^{n-d}) \triangleq \prod_{i=1}^{n} p(y_{i}|y^{i-1}, x^{i-d})$$

a. Prove that

i.

$$I(X^n \to Y^n) \triangleq E\left[\log \frac{p(y^n||x^n)}{p(y^n)}\right]$$

ii.

$$p(y^n, x^n) = p(y^n || x^n) p(x^n || y^{n-1})$$

iii.

$$I(X^n \to Y^n) \ge 0$$

and equals zero if and only if  $p(y^n || x^n) = p(y^n)$ 

iv.

$$I(X^n; Y^n) = I(X^n \to Y^n) + I(0Y^{n-1} \to X^n),$$

where the term ,  $\emptyset Y^{n-1}$ , denotes the concatenation of 0 (Null) to the sequence  $Y^{N-1}$ , i.e.  $(\emptyset, Y_1, Y_2, \dots Y_{n-1})$ .

b. Prove that in general

$$I(X^n \to Y^n) \le I(X^n; Y^n)$$

and equality holds if and only if  $p(x^n||y^{n-1}) = p(x^n)$ .

c. Suggest (without proof) properties similar to those of mutual information that should hold for directed information.

Solution: Directed Information and causal conditioning.

(a) i.

$$I(X^n \to Y^n) = H(Y^n) - H(Y^n||X^n)$$
  
=  $E\left[\log \frac{1}{P(y^n)}\right] - E\left[-\log P(y^n||x^n)\right]$   
=  $E\left[\log \frac{P(y^n||x^n)}{P(y^n)}\right]$ 

ii.

$$p(y^{n}, x^{n}) = \prod_{i=1}^{n} p(y_{i}, x_{i} | y^{i-1}, x^{i-1})$$
  
= 
$$\prod_{i=1}^{n} p(x_{i} | y^{i-1}, x^{i-1}) p(y_{i} | y^{i-1}, x^{i})$$
  
= 
$$p(x^{n} | | y^{n-1}) p(y^{n} | | x^{n})$$

iii. Directed information is a sum of non-negative terms, i.e.  $I(X^i; Y_i | Y^{i-1}) \ge 0$  and therefore it is nonnegative. From the definition we see that directed information equals zero if and only if  $p(y^n | | x^n) = p(y^n)$  for all  $(x^n, y^n)$ .

iv.

$$I(X^{n}; Y^{n}) = E\left[\log \frac{p(Y^{n}, X^{n})}{p(Y^{n})p(X^{n})}\right]$$
  
=  $E\left[\log \frac{p(Y^{n}||X^{n})p(X^{n}||X^{n-1})}{p(Y^{n})p(X^{n})}\right]$   
=  $E\left[\log \frac{p(Y^{n}||X^{n})}{p(Y^{n})}\right] + E\left[\log \frac{p(X^{n}||Y^{n-1})}{p(X^{n})}\right]$   
=  $I(X^{n} \to Y^{n}) + I(Y^{n-1} \to X^{n})$ 

(b) The inequality  $I(X^n \to Y^n) \leq I(X^n; Y^n)$  follows immediately from the exercise in (a) and equality holds if and only if  $p(x^n||y^{n-1}) = p(x^n)$ . This condition is equivalent to the condition that feedback is not used in the encoding scheme.