

Lecture 1: Method of types and strong typicality

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I. A TYPES: DEFINITION AND PROPERTIES

The method of types evolved from notions of strong typicality; some of the ideas were used by Wolfowitz [4] to prove channel capacity theorems. The method was fully developed by Csiszar and Korner [1], who derived the main theorems of information theory from this viewpoint.

We will start the lecture by defining a type of a sequence. Let $x^n = (x_1, x_2, \dots, x_n)$ be a sequence from alphabet $\mathcal{X} = (a_1, a_2, a_3, \dots, a_{|\mathcal{X}|})$. Let $N(a|x^n)$ be the number of times that a appears in sequence x^n .

Definition 1 (Type) The type P_{x^n} (or empirical probability distribution) of a sequence x^n is the relative proportion of occurrences of each symbol of \mathcal{X} , i.e., $P_{x^n}(a) = \frac{N(a|x^n)}{n}$ for all $a \in \mathcal{X}$.

Example 1 Let $\mathcal{X} = \{0, 1, 2\}$, let $n = 5$ and $x^5 = (1, 1, 2, 2, 0)$. Then $N(0|x^5) = 1$, $N(1|x^5) = 2$ and $N(2|x^5) = 2$. Hence, $P_{x^5} = (\frac{1}{5}, \frac{2}{5}, \frac{2}{5})$.

Definition 2 (all possible types) Let \mathcal{P}_n be the collection of all possible types of sequences of length n .

For example, if $\mathcal{X} = \{0, 1\}$, the set of possible types with denominator n is

$$\mathcal{P}_n = \left\{ (P(0), P(1)) : \left(\frac{0}{n}, \frac{n}{n} \right), \left(\frac{1}{n}, \frac{n-1}{n} \right), \dots, \left(\frac{n}{n}, \frac{0}{n} \right) \right\}. \quad (1)$$

Lemma 1 An upper bound for $|\mathcal{P}_n|$:

$$|\mathcal{P}_n| \leq (n+1)^{|\mathcal{X}|}. \quad (2)$$

Proof: There are $|\mathcal{X}|$ components in the vector that specifies P_{x^n} . The numerator in each component can take on only $n + 1$ values. So there are at most $(n + 1)^{|\mathcal{X}|}$ choices for the type vector. ■

Definition 3 (Type class) Let $P \in \mathcal{P}_n$, The set of sequences of length n with type P is called type class of P , denoted $T(P)$:

$$T(P) = \{x^n : P_{x^n} = P\} \quad (3)$$

Lets us now define the notation $Q^n(x^n)$ that emphasis that X^n is distributed i.i.d according to $Q(x)$. In other words,

$$Q^n(x^n) \triangleq \prod_{i=1}^n Q(x_i). \quad (4)$$

Theorem 1 (Probability of a sequence in the type class) If $X \sim Q$ i.i.d., the probability of x^n depends only on the type of x^n , i.e., P_{x^n}

$$Q^n(x^n) = 2^{-n(H(P_{x^n}) + D(P_{x^n}||Q))} \quad (5)$$

Proof: Consider

$$\log Q^n(x^n) = \sum_{i=1}^n \log Q(x_i) \quad (6)$$

$$\stackrel{(a)}{=} \sum_{a \in \mathcal{X}} N(a|x^n) \log Q(a) \quad (7)$$

$$\stackrel{(b)}{=} n \sum_{a \in \mathcal{X}} P_{x^n}(a) \log Q(a) \quad (8)$$

$$= n \sum_{a \in \mathcal{X}} P_{x^n}(a) \log \frac{Q(a)}{P_{x^n}(a)} \cdot P_{x^n}(a) \quad (9)$$

$$= n(-H(P) - D(P||Q)), \quad (10)$$

where

(a) follows because each $a \in \mathcal{X}$ contributes exactly $\log Q(a)$ times it's number of occurences in x^n to the sum in (6).

(b) follows from the definition of $P_{x^n}(a)$.

Hence we obtained

$$Q^n(x^n) = 2^{(-nH(P)+D(P||Q))}. \quad (11)$$

■

Corollary 1 if x^n is in the type class of Q , then we get $Q^n(x^n) = 2^{-nH(P_{x^n})}$.

The following theorem tells us how many sequences, asymptotically, exist of type $P \in \mathcal{P}_n$.

Theorem 2 (size of a type class) For any type $P \in \mathcal{P}_n$

$$|T(p)| \doteq 2^{nH(P)} \quad (12)$$

Where $a_n \doteq b_n$ if $\lim_{n \rightarrow \infty} \frac{1}{n} \log\left(\frac{a_n}{b_n}\right) = 0$.

Example 2 Question: How many binary sequences of length n with 50% 0 and 50% 1 exists?

Answer: An exact calculation yields $\binom{n}{\frac{n}{2}}$. An asymptotic calculation Using Theorem 2 yields that $\binom{n}{\frac{n}{2}} \doteq 2^n$.

There are two possible ways to prove Theorem 2, one is a combinatorial proof and the other is a probabilistic. We will provide both proofs in two different subsections.

II. COMBINATORIAL PROOF OF THEOREM 2

Lemma 2 (Stirling's formula) :

The combinatorial proof is based on Stirling's Formula:

$$\sqrt{2\pi n} \left(\frac{n}{e}\right)^n \leq n! \leq \sqrt{2\pi n} \left(\frac{n}{e}\right)^n e^{\frac{1}{12n}} \quad (13)$$

proof of Theorem 2:

$$|T(P)| = \binom{n}{nP(a_1), nP(a_2), \dots, nP(a_X)} = \frac{n!}{(nP(a_1))!(nP(a_2))! \dots (nP(a_X))!} \quad (14)$$

Using Stirling's formula with equation (5) we get:

$$n! \doteq \left(\frac{n}{e}\right)^n \quad (15)$$

a_1	a_2	\dots	$a_{ \mathcal{X} }$
$nP(a_1)$	$nP(a_2)$		$nP(a_{ \mathcal{X} })$

Fig. 1. The total length of the sequence is n and the part of the sequence that equals to a_i is $nP(a_i)$

$$|T(P)| \doteq \frac{n^n}{(nP(a_1))^{nP(a_1)}(nP(a_2))^{nP(a_2)} \dots (nP(a_{|\mathcal{X}|}))^{nP(a_{|\mathcal{X}|})}} \quad (16)$$

$$= \frac{n^n}{(n)^{nP(a_1)}(n)^{nP(a_2)} \dots (n)^{nP(a_{|\mathcal{X}|})} \prod_{i=1}^{|\mathcal{X}|} P(a_i)^{nP(a_i)}} \quad (17)$$

$$= \frac{1}{\prod_{i=1}^{|\mathcal{X}|} P(a_i)^{nP(a_i)}} \quad (18)$$

Hence:

$$|T(P)| = 2^{\log |T(P)|} \doteq 2^{-n \sum_{i=1}^{|\mathcal{X}|} P(a_i) \log(P(a_i))} = 2^{nH(P)} \quad (19)$$

■

III. PROBABILISTIC PROOF OF THEOREM 2:

From the probabilistic proof we will obtain two bounds that implies Theorem 2. The bounds are:

$$\frac{2^{nH(P)}}{(n+1)^{|\mathcal{X}|}} \leq |T(P)| \leq 2^{nH(P)}. \quad (20)$$

Proof: Let's assume the sequence X^n is distributed i.i.d according to $P(x)$. Now consider the following:

$$1 \geq \Pr(x^n \in T(P)) \quad (21)$$

$$\stackrel{(a)}{=} \sum_{x^n \in T(P)} \Pr(x^n) \quad (22)$$

$$\stackrel{(b)}{=} \sum_{x^n \in T(P)} 2^{-nH(P)} \quad (23)$$

$$= |T(P)| 2^{-nH(P)}. \quad (24)$$

Equality (a) follows from the fact that the probability of a subset equals to the sum of the probabilities of each element in the subset. For example, if we have a set A, B, C where

the probability of choosing A is P_A , B is P_B and C is P_C , where $P_A + P_B + P_C = 1$, then the probability of choosing from the subset (A, B) is $P_A + P_B$. Equality (b) follows from Theorem 1.

Therefore:

$$|T(P)| \leq 2^{nH(P)} \quad (25)$$

In order to prove the other part we need the following lemma:

Lemma 3 $P^n(T(P)) \geq P^n(T(Q))$

Proof: Let X^n be of a type P . The term $P^n(T(P))$ is the probability of type class $T(P)$ where the sequences of length n are drawn according to $P(x^n) = \prod_{i=1}^n P(x_i)$, and let $Q \in \mathcal{P}_n$.

Consider

$$\frac{P^n(T(P))}{P^n(T(Q))} \stackrel{(a)}{=} \frac{|T(P)| \prod_{a \in \mathcal{X}} P(a)^{nP(a)}}{|T(Q)| \prod_{a \in \mathcal{X}} P(a)^{nQ(a)}} \quad (26)$$

$$\stackrel{(b)}{=} \frac{\binom{n}{nP(a_1), nP(a_2), \dots, nP(a_{|\mathcal{X}|})} \prod_{a \in \mathcal{X}} P(a)^{nP(a)}}{\binom{n}{nQ(a_1), nQ(a_2), \dots, nQ(a_{|\mathcal{X}|})} \prod_{a \in \mathcal{X}} P(a)^{nQ(a)}} \quad (27)$$

$$\stackrel{(c)}{=} \prod_{a \in \mathcal{X}} \frac{(nQ(a))!}{(nP(a))!} P(a)^{n(P(a)-Q(a))} \quad (28)$$

(a) Using the fact that probability of each type $P_{x^n} \in \mathcal{P}_n$ is given by:

$$P_{x^n} = \prod_{i=1}^n P(x_i) = \prod_{a \in \mathcal{X}} P(a)^{N(a|x^n)} = \prod_{a \in \mathcal{X}} P(a)^{nP(a)}.$$

(b) Using combinatorial math it is known that the number of possibilities to arrange a

vector $\{x^n : P_{x^n} = P\}$ is: $\binom{n}{nP(a_1), nP(a_2), \dots, nP(a_{|\mathcal{X}|})}$.

$$(c) \frac{\binom{n}{nP(a_1), nP(a_2), \dots, nP(a_{|\mathcal{X}|})}}{\binom{n}{nQ(a_1), nQ(a_2), \dots, nQ(a_{|\mathcal{X}|})}} = \prod_{a \in \mathcal{X}} \frac{(nQ(a))!}{(nP(a))!}$$

Using the simple bound $\frac{m!}{n!} \geq n^{m-n}$ we obtain:

$$\frac{P^n(T(P))}{P^n(T(Q))} \geq \prod_{a \in \mathcal{X}} (nP(a))^{nQ(a)-nP(a)} P(a)^{n(P(a)-Q(a))} \quad (29)$$

$$= \prod_{a \in \mathcal{X}} n^{n(Q(a)-P(a))} \quad (30)$$

$$= n^{n(\sum_{a \in \mathcal{X}} Q(a) - \sum_{a \in \mathcal{X}} P(a))} \quad (31)$$

$$= n^{n(1-1)} = 1 \quad (32)$$

■

Using Lemma 3 let us show that $|T(P)| \geq \frac{2^{nH(P)}}{(n+1)^{|\mathcal{X}|}}$:

$$1 = \sum_{Q \in \mathcal{P}_n} P^n(T(Q)) \quad (33)$$

$$\leq \sum_{Q \in \mathcal{P}_n} \max_Q P^n(T(Q)) \quad (34)$$

$$\stackrel{(a)}{=} \sum_{Q \in \mathcal{P}_n} P^n(T(P)) \quad (35)$$

$$\stackrel{(b)}{\leq} (n+1)^{|\mathcal{X}|} P^n(T(P)) \quad (36)$$

$$\stackrel{(c)}{=} (n+1)^{|\mathcal{X}|} \sum_{x^n \in T(P)} 2^{-nH(P)} \quad (37)$$

$$= (n+1)^{|\mathcal{X}|} |T(P)| 2^{-nH(P)} \quad (38)$$

(a) Using theorem 2 it is clear that: $\max_Q P^n(T(Q)) = P^n(T(P))$.

(b) Using Lemma 1.

(c) Using Theorem 3.

Therefore our final result is:

$$\frac{2^{nH(P)}}{(n+1)^{|\mathcal{X}|}} \leq |T(P)| \leq 2^{nH(P)} \quad (39)$$

which implies that:

$$|T(P)| \doteq 2^{nH(P)} \quad (40)$$

■

IV. PROBABILITY OF A TYPE AND OF A SET OF TYPES (SANOV'S THEOREM)

Theorem 3 The probability of the type class $T(P)$ where the sequences are drawn i.i.d. $\sim Q$ is

$$Q^n(T(P)) \doteq 2^{-n(D(P)||Q)}. \quad (41)$$

Proof:

$$Q^n(T(P)) = \sum_{x^n \in T(P)} Q(x^n) \quad (42)$$

$$\stackrel{(a)}{=} \sum_{x^n \in T(P)} 2^{-n(H(P_{x^n}) + D(P_{x^n} \| Q))} \quad (43)$$

$$\stackrel{(b)}{=} \sum_{x^n \in T(P)} 2^{-n(H(P) + D(P \| Q))} \quad (44)$$

$$= |T(P)| 2^{-n(H(P) + D(P \| Q))} \quad (45)$$

$$\stackrel{(c)}{=} 2^{-nD(P \| Q)}, \quad (46)$$

where (a) follows from Theorem 1, (b) from the fact that all sequences have the same type $P_{x^n} = P$ and (c) from Theorem 2.

One can also obtain more explicit bounds by using the explicit bounds on $|T(P)|$ given in (39):

$$\frac{2^{-nD(P \| Q)}}{(n+1)^{|\mathcal{X}|}} \leq Q^n(T(P)) \leq 2^{-nD(P \| Q)} \quad (47)$$

■

Next we state Sanov's theorem, [3] which was generalized by Csiszar [2] using the method of types. It also opened a new field in statistics called *Large Deviation*.

Theorem 4 (Sanov's Theorem) Let $X \sim Q$ i.i.d. and let E be a set of probabilities that is the closure of its interior, then:

$$\lim_{n \rightarrow \infty} \log Q^n(E) = -\min_{P \in E} D(P \| Q) = -D(P^* \| Q), \quad (48)$$

where $Q^n(E)$ is the probability that $x^n \in E$ i.e. $Q^n(E) = \Pr(P \in E)$ and P^* is defined as $P^* = \arg \min_{P \in E} D(P \| Q)$.

To get more intuitive understanding we can think of $D(P^* \| Q)$ as the minimum distance between E space and Q as shown in the figure:

$$Q^n(E) \doteq 2^{-nD(P^* \| Q)} \quad (49)$$

$$P^* = \arg \min_{P \in E} D(P \| Q) \quad (50)$$

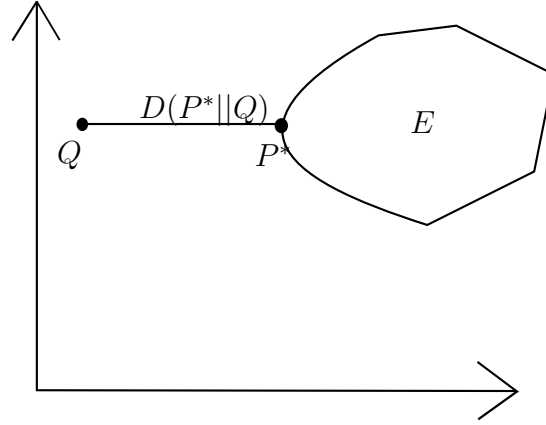


Fig. 2. Let $X \sim Q$ then P^* is the type $P \in E$ that gives the minimum to $D(P||Q)$.

Example 3 Let $Q(x = 1) = Q(x = -1) = \frac{1}{2}$, What is the probability of getting an empirical distribution that satisfies: $P(x = 1) \geq 0.8$, $P(x = -1) \leq 0.2$?

Answer: P^* is the probability $P(x = 1) = 0.8$, $P(x = -1) = 0.2$ so by using Sanov theorem and Theorem 4 we get our result: $Q(E) \doteq 2^{-nD(P^*||Q)}$

Proof of Theorem 4: First we will find the upper bound

$$Q^n(E) = \sum_{P \in E \cap \mathcal{P}_n} Q^n(T(P)) \quad (51)$$

$$\stackrel{(a)}{\leq} \sum_{P \in E \cap \mathcal{P}_n} 2^{-nD(P||Q)} \quad (52)$$

$$\leq \sum_{P \in E \cap \mathcal{P}_n} \max_{p \in E \cap \mathcal{P}_n} 2^{-nD(p||Q)} \quad (53)$$

$$= \sum_{P \in E \cap \mathcal{P}_n} 2^{-n \min_{P \in E \cap \mathcal{P}_n} D(P||Q)} \quad (54)$$

$$\stackrel{(b)}{\leq} (n+1)^{|\mathcal{X}|} 2^{-n \min_{P \in E \cap \mathcal{P}_n} D(P||Q)}, \quad (55)$$

where (a) follows from Theorem 3, and (b) follows from the fact that $|E| \leq |\mathcal{P}_n|$ and the bound on the number of types (Lemma 1).

The minimum of $\min_{P \in E \cap \mathcal{P}_n} D(P||Q)$ exists since E is closed, further more because E is the closure of its interior and Divergence is continuous, it implies that the $\lim_{n \rightarrow \infty} \min_{P \in E \cap \mathcal{P}_n} D(P||Q)$ exists and is obtained by some $P^* \in E$.

Now we will find the lower bound:

$$Q^n(E) = \sum_{P \in E \cap \mathcal{P}_n} Q^n(T(P)) \quad (56)$$

$$\stackrel{(a)}{\geq} \min_{P \in E \cap \mathcal{P}_n} Q(T(P)) \quad (57)$$

$$\stackrel{(b)}{=} \min_{P \in E \cap \mathcal{P}_n} 2^{-nD(P||Q)} \quad (58)$$

where (a) follows from the fact that we take into consideration only one type and (b) According to Theorem 3.

We can obtain a more explicit bound using (47) in the last step:

$$Q^n(E) \geq \frac{1}{(n+1)^{|\mathcal{X}|}} 2^{-n \min_{P \in E \cap \mathcal{P}_n} D(P||Q)}. \quad (59)$$

Combining the lower bound (55) and upper bound (59) we have

$$\frac{1}{(n+1)^{|\mathcal{X}|}} 2^{-n \min_{P \in E \cap \mathcal{P}_n} D(P||Q)} \leq Q^n(E) \leq (n+1)^{|\mathcal{X}|} 2^{-n \min_{P \in E \cap \mathcal{P}_n} D(P||Q)}. \quad (60)$$

$$(61)$$

which implies

$$Q^n(E) \doteq 2^{-nD(P^*||Q)} \quad (62)$$

■

V. JOINT TYPE

Definition 4 (Joint type) The type P_{x^n, y^n} (or empirical probability distribution) of a pair-sequence (x^n, y^n) is the relative proportion of occurrences of each pair-symbol of $\mathcal{X} \times \mathcal{Y}$, i.e., $P_{x^n, y^n}(a, b) = \frac{N(a, b|x^n, y^n)}{n}$ for all $a \in \mathcal{X}$ and $b \in \mathcal{Y}$.

Example 4 Let $\mathcal{X} = \{0, 1\}$, and $\mathcal{Y} = \{A, B\}$. let $n = 5$ and $x^5 = (1, 1, 0, 1, 0)$ and $y^5 = (A, A, B, A, B)$. Then $N(0, A|x^5) = 0$, $N(0, B|x^5) = 2$, $N(1, A|x^5) = 3$ and $N(1, B|x^5) = 0$.

Theorem 5 (Conditional type)

Let us define the *conditional type* $P_{x^n|y^n}$ (or conditional empirical distribution)

$$P_{x^n|y^n}(a|b) \triangleq \frac{N((a, b)|x^n, y^n)}{N(b|y^n)} \quad (63)$$

$$= \frac{P_{X^n, Y^n}(a, b)}{P_{Y^n}(b)}. \quad (64)$$

Let $W(y|x) \in \mathcal{P}^n(x|y)$ be a conational probability, The conditional type $T_W(y^n)$

$$T_W(y^n) = \{x^n \in \mathcal{X}^n : P_{X^n|Y^n}(a|b) = W_{X|Y}(a|b), \forall a, b \in \mathcal{X}, \mathcal{Y}\} \quad (65)$$

$$= \{x^n \in \mathcal{X}^n : P_{X^n, Y^n}(a, b) = W_{X|Y}(a|b)P_{Y^n}(b), \forall a, b \in \mathcal{X}, \mathcal{Y}\} \quad (66)$$

$$H(X|Y) = - \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} P(x, y) \log P(x|y) \quad (67)$$

$$P_{X, Y}(a, b) = P_{Y^n}(b)W_{X|Y}(a|b) \quad (68)$$

Than:

$$|T_W(y^n)| \doteq 2^{nH(X|Y)} \quad (69)$$

Proof:

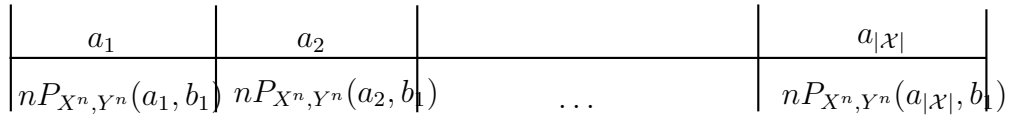
b_1	b_2	\dots	$b_{ \mathcal{Y} }$
$nP_{Y^n}(b_1)$	$nP_{Y^n}(b_2)$	\dots	$nP_{Y^n}(b_{ \mathcal{Y} })$

Fig. 3. Length of each b_i .

Now if we have b_1 we get:

Therefore we can use combinatorial proof as we did in the non conditional case:

$$\left(nP_{x^n, y^n}(a_1, b_1) nP_{x^n, y^n}(a_2, b_1) \dots nP_{x^n, y^n}(a_{|\mathcal{X}|}, b_1) \right) \doteq 2^{nH(X|y=b_1)P_{y^n}(b_1)} \quad (70)$$

Fig. 4. Length of each a_i given b_1 .

$$\left(nP_{Y^n}(b_1)P_{x^n|y^n}(a_1|b_1)nP_{Y^n}(b_1)P_{x^n|y^n}(a_2|b_1)\dots nP_{Y^n}(b_1)P_{x^n|y^n}(a_{|\mathcal{X}|}|b_1) \right) \doteq 2^{nP_{Y^n}(b_1)H(X|y=b_1)} \quad (71)$$

$$|T_W(y^n)| \doteq \prod_{i=1}^{|\mathcal{Y}|} 2^{nH(X|y=b_i)P_{Y^n}(b_i)} = 2^{nH(X|Y)} \quad (72)$$

VI. STRONG TYPICALITY

Definition 5 (ϵ -strongly typical) A sequence $x^n \in \mathcal{X}$ is said to be ϵ -strongly typical with respect to a distribution $P(x)$ on \mathcal{X} if

- 1) For all $a \in \mathcal{X}$ with $P_X(a) > 0$ we have

$$|P_{x^n}(a) - P_X(a)| \leq \frac{\epsilon}{|\mathcal{X}|} \quad (73)$$

- 2) If $P_X(a) = 0$ then $P_{x^n}(a) = 0$.

Definition 6 (ϵ -joint strongly typical) A pair of sequences $(x^n, y^n) \in \mathcal{X} \times \mathcal{Y}$ is said to be ϵ -joint strongly typical with respect to a distribution $P(x, y)$ on $\mathcal{X} \times \mathcal{Y}$ if

- 1) For all $(a, b) \in \mathcal{X} \times \mathcal{Y}$ with $P_{X,Y}(a, b) > 0$ we have

$$|P_{x^n, y^n}(a, b) - P_{X,Y}(a, b)| \leq \frac{\epsilon}{|\mathcal{X}||\mathcal{Y}|} \quad (74)$$

- 2) If $P_{X,Y}(a, b) > 0$ then $P_{x^n, y^n}(a, b) = 0$.

Definition 7 (Strongly typical set) The set of sequences $(x^n, y^n) \in \mathcal{X}^n \times \mathcal{Y}^n$ that are ϵ -joint strongly typical is called *strongly typical set* and is denoted as $T_\epsilon^{(n)}(X, Y)$ or $T_\epsilon^{(n)}(P_{X,Y})$. I.e.,

$$T_\epsilon^{(n)}(X, Y) \triangleq \{(x^n, y^n) \in \mathcal{X}^n \times \mathcal{Y}^n : |P_{x^n, y^n} - P(x, y)| \leq \epsilon\} \quad (75)$$

In a shorter notation we write it as

$$T_\epsilon^{(n)}(X, Y) \triangleq \{x^n, y^n : |P_{x^n, y^n} - P(x, y)| \leq \epsilon\}. \quad (76)$$

Definition 8 (Strongly conditional typical set $T_\epsilon^{(n)}(X|y^n)$)

$$\begin{aligned} T_\epsilon^{(n)}(Y|x^n) &\triangleq \{y^n : (x^n, y^n) \in T_\epsilon^{(n)}(X, Y)\} \\ &= \{y^n : |P_{x^n, y^n} - P(x, y)| \leq \epsilon\}. \end{aligned} \quad (77)$$

The following lemma follows directly from the Sanov's Theorem.

Lemma 4 Suppose w.l.o.g. that $Q(x) > 0$ for all $x \in \mathcal{X}$ (otherwise shrink the alphabet to its effective size) and that X_i are i.i.d. $\sim Q(x)$. Then there exists $\epsilon'(\epsilon)$ such that $\epsilon'(\epsilon) \rightarrow 0$ as $\epsilon \rightarrow 0$ such that for all sufficiently large n

$$2^{-n[D(P||Q)+\epsilon']} \leq \Pr(X^n \in T_\epsilon^{(n)}(P)) \leq 2^{-n[D(P||Q)-\epsilon']}. \quad (78)$$

Lemma 5 Consider a joint distribution $P_{X,Y}$ with marginal P_X and P_Y . Generate X^n i.i.d. $\sim P_X$ and $Y^n \sim P_Y$. Then

$$2^{-n[I(X;Y)+\epsilon']} \leq \Pr((X^n, Y^n) \in T_\epsilon^{(n)}(X, Y)) \leq 2^{-n[I(X;Y)-\epsilon']}. \quad (79)$$

VII. ALTERNATIVE PROOFS OF PROPERTIES OF STRONGLY TYPICAL SET BASED ONLY ON PROBABILISTIC METHODS

Let us define the strong type slightly differen but equivalently as follows:

$$T_\epsilon^{(n)}(P_X) = T_\epsilon^{(n)}(X) = \{x^n \in \mathcal{X}^n : |P_{x^n}(a) - P_X(a)| \leq \epsilon P_X(a)\} \quad (80)$$

and

$$T_\epsilon^{(n)}(P_{XY}) = T_\epsilon^{(n)}(X, Y) = \{x^n, y^n : |P_{x^n y^n}(a, b) - P_{XY}(a, b)| \leq \epsilon P_{XY}(a, b)\}. \quad (81)$$

Properties:

1) If $X^n \sim P_X$ i.i.d. , then

$$\lim_{n \rightarrow \infty} \Pr \{X^n \in T_\epsilon^{(n)}(X)\} = 1 \quad (82)$$

or equivalently

$$1 - \delta_n(\epsilon) \leq \Pr \{X^n \in T_\epsilon^{(n)}(X)\} \leq 1, \quad (83)$$

where $\forall \epsilon > 0, \lim_{n \rightarrow \infty} \delta_n(\epsilon) = 0$.

Proof: Because of the L.L.N. ■

2) For $X^n \sim P_X$, i.i.d. , then for any sequence $x^n \in T_\epsilon^{(n)}(X)$ we have

$$2^{-nH(X)(1+\epsilon)} \stackrel{(i)}{\leq} p(x^n) = \prod_{i=1}^n P_X(x_i) \stackrel{(ii)}{\leq} 2^{-nH(X)(1-\epsilon)}. \quad (84)$$

Proof: Note that $p(x^n) = \prod_{a \in \mathcal{X}} P_X(a)^{N(a|x^n)}$, hence, the left hand side of the inequality (i) follows from

$$\frac{1}{n} \log p(x^n) = \frac{1}{n} \sum_{a \in \mathcal{X}} N(a|x^n) \log P_X(a) \quad (85)$$

$$= \sum_{a \in \mathcal{X}} P_{x^n}(a) \log P_X(a) \quad (86)$$

$$\stackrel{(a)}{\geq} \sum_{a \in \mathcal{X}} (1 + \epsilon) P_X(a) \log P_X(a) \quad (87)$$

$$= -H(X)(1 + \epsilon), \quad (88)$$

where step (a) follows from (80).

The right hand side of the inequality (ii) follows from

$$\frac{1}{n} \log p(x^n) = \frac{1}{n} \sum_{a \in \mathcal{X}} N(a|x^n) \log P_X(a) \quad (89)$$

$$= \sum_{a \in \mathcal{X}} P_{x^n}(a) \log P_X(a) \quad (90)$$

$$\leq \sum_{a \in \mathcal{X}} (1 - \epsilon) P_X(a) \log P_X(a) \quad (91)$$

$$= -H(X)(1 - \epsilon). \quad (92)$$

■

3) The size of the strongly typical set can be bounded as

$$(1 - \delta_{\epsilon,n}) 2^{nH(X)(1-\epsilon)} \stackrel{(i)}{\leq} |T_\epsilon^{(n)}(X)| \stackrel{(ii)}{\leq} 2^{nH(X)(1+\epsilon)}. \quad (93)$$

Proof: Recall that under the assumption that $X^n \sim P_X$, i.i.d. , we have $(1 - \delta) \leq \Pr \{X^n \in T_\epsilon^{(n)}(X)\} \leq 1$. Now we prove the left hand side inequality (i):

$$\Pr \{X^n \in T_\epsilon^{(n)}(X)\} = \sum_{x^n \in T_\epsilon^{(n)}(X)} p(x^n) \quad (94)$$

$$\leq \sum_{x^n \in T_\epsilon^{(n)}(X)} 2^{-n(H(X)(1-\epsilon))} \quad (95)$$

$$= |T_\epsilon^{(n)}(X)| 2^{-nH(X)(1-\epsilon)}, \quad (96)$$

and because $(1 - \delta) \leq \Pr \{X^n \in T_\epsilon^{(n)}(X)\}$, we get that $(1 - \delta)2^{nH(X)(1-\epsilon)} \leq |T_\epsilon^{(n)}(X)|$. The right hand side inequality (ii) follows from

$$1 \geq \Pr \{X^n \in T_\epsilon^{(n)}(X)\} \quad (97)$$

$$= \sum_{x^n \in T_\epsilon^{(n)}(X)} p(x^n) \quad (98)$$

$$\geq \sum_{x^n \in T_\epsilon^{(n)}(X)} 2^{-nH(X)(1+\epsilon)} \quad (99)$$

$$= |T_\epsilon^{(n)}(X)| 2^{-nH(X)(1+\epsilon)}, \quad (100)$$

therefore, $|T_\epsilon^{(n)}(X)| \leq 2^{nH(X)(1+\epsilon)}$. ■

Conditionally strong typical set

For a given $x^n \in \mathcal{X}^n$, let us define

$$T_\epsilon^{(n)}(Y|x^n) = \{y^n : (x^n, y^n) \in T_\epsilon^{(n)}(X, Y)\}. \quad (101)$$

Notice that if $(x^n, y^n) \in T_\epsilon^{(n)}(X, Y)$, then surely $x^n \in T_\epsilon^{(n)}(X)$.

Properties:

- 1) If $x^n \in T_{\epsilon_x}^{(n)}(X)$, $p(y^n|x^n) = \prod_{i=1}^n P_{Y|X}(y_i|x_i)$ (DMC), then for all $0 < \epsilon_x < \epsilon$

$$1 - \delta_{\epsilon, \epsilon_x, n} \leq \Pr \{y^n \in T_\epsilon^{(n)}(Y|x^n)\} \leq 1 \quad (102)$$

where $\delta_{\epsilon, \epsilon_x, n} \rightarrow 0$ as $n \rightarrow \infty$ for all $0 < \epsilon_x \leq \epsilon$.

Proof: Follows directly from the L.L.N. ■

- 2) If $y^n \in T_\epsilon^{(n)}(Y|x^n)$, $p(y^n|x^n) = \prod_{i=1}^n P_{Y|X}(y_i|x_i)$ (DMC), then

$$2^{-nH(Y|X)(1+\epsilon)} \stackrel{(i)}{\leq} p(y^n|x^n) = \prod_{i=1}^n P_{Y|X}(y_i|x_i) \stackrel{(ii)}{\leq} 2^{-nH(Y|X)(1-\epsilon)}. \quad (103)$$

Proof: Notice that

$$p(y^n|x^n) = \prod_{i=1}^n P_{Y|X}(y_i|x_i) = \prod_{\substack{a \in \mathcal{X} \\ b \in \mathcal{Y}}} P_{Y|X}(b|a)^{N(a,b|x^n,y^n)}. \quad (104)$$

Now, the left hand side of the inequality (i) follows from

$$\frac{1}{n} \log p(y^n|x^n) = \frac{1}{n} \sum_{\substack{a \in \mathcal{X} \\ b \in \mathcal{Y}}} N(a,b|x^n,y^n) \log P_{Y|X}(b|a) \quad (105)$$

$$= \sum_{\substack{a \in \mathcal{X} \\ b \in \mathcal{Y}}} P_{x^n,y^n}(a,b) \log P_{Y|X}(b|a) \quad (106)$$

$$\geq \sum_{\substack{a \in \mathcal{X} \\ b \in \mathcal{Y}}} (1 + \epsilon) P_{XY}(a,b) \log P_{Y|X}(b|a) \quad (107)$$

$$= -H(Y|X)(1 + \epsilon). \quad (108)$$

The right hand side of the inequality (ii) follows from

$$\frac{1}{n} \log p(y^n|x^n) = \frac{1}{n} \sum_{\substack{a \in \mathcal{X} \\ b \in \mathcal{Y}}} N(a,b|x^n,y^n) \log P_{Y|X}(b|a) \quad (109)$$

$$= \sum_{\substack{a \in \mathcal{X} \\ b \in \mathcal{Y}}} P_{x^n,y^n}(a,b) \log P_{Y|X}(b|a) \quad (110)$$

$$\leq \sum_{\substack{a \in \mathcal{X} \\ b \in \mathcal{Y}}} (1 - \epsilon) P_{XY}(a,b) \log P_{Y|X}(b|a) \quad (111)$$

$$= -H(Y|X)(1 - \epsilon). \quad (112)$$

■

3) Given $x^n \in T_{\epsilon_x}^{(n)}(X)$, then

$$(1 - \delta_{\epsilon,\epsilon_x,n}) 2^{nH(Y|X)(1+\epsilon)} \stackrel{(i)}{\leq} |T_{\epsilon}^{(n)}(Y|x^n)| \stackrel{(ii)}{\leq} 2^{nH(Y|X)(1-\epsilon)}. \quad (113)$$

The proof is done in a similar way to the proof of (93).

Lemma 6 Consider a joint PMF P_{XY} . Let $x^n \in T_{\epsilon_x}^{(n)}(X)$ and y^n drawn i.i.d. according to P_Y and independent of x^n , then

$$2^{-n(I(X;Y)+\delta_\epsilon)} \stackrel{(i)}{\leq} \Pr \{Y^n \in T_{\epsilon}^{(n)}(Y|x^n)\} \stackrel{(ii)}{\leq} 2^{-n(I(X;Y)-\delta_\epsilon)} \quad (114)$$

for $\delta_\epsilon \rightarrow 0$ as $\epsilon \rightarrow 0$.

Proof: The left hand side inequality (i) follows from

$$\Pr \{Y^n \in T_\epsilon^{(n)}(Y|x^n)\} = \sum_{y^n \in T_\epsilon^{(n)}(Y|x^n)} p(y^n) \quad (115)$$

$$\geq \sum_{y^n \in T_\epsilon^{(n)}(Y|x^n)} 2^{-nH(Y)(1+\epsilon)} \quad (116)$$

$$= |T_\epsilon^{(n)}(Y|x^n)| 2^{-nH(Y)(1+\epsilon)} \quad (117)$$

$$\geq (1 - \delta_{\epsilon,n}) 2^{n(H(Y|X)(1-\epsilon))} 2^{-nH(Y)(1+\epsilon)} \quad (118)$$

$$= (1 - \delta_{\epsilon,n}) 2^{-n(I(X;Y)+\delta_{\epsilon,n})}. \quad (119)$$

The right hand side inequality (ii) follows from

$$\Pr \{Y^n \in T_\epsilon^{(n)}(Y|x^n)\} = \sum_{y^n \in T_\epsilon^{(n)}(Y|x^n)} p(y^n) \quad (120)$$

$$\leq \sum_{y^n \in T_\epsilon^{(n)}(Y|x^n)} 2^{-nH(Y)(1-\epsilon)} \quad (121)$$

$$= |T_\epsilon^{(n)}(Y|x^n)| 2^{-nH(Y)(1-\epsilon)} \quad (122)$$

$$\leq 2^{n(H(Y|X)(1+\epsilon))} 2^{-nH(Y)(1-\epsilon)} \quad (123)$$

$$= 2^{-n(I(X;Y)-\delta_\epsilon)}. \quad (124)$$

■

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