Final Exam - Moed A

Total time for the exam: 3 hours!

Please copy the following sentence and sign it: " I am respecting the rules of the exam: Signature:_____ "

Important: For **True / False** questions, copy the statement to your notebook and write clearly true or false. You should prove the statement if true, or disprove it, e.g. by providing a counter-example, otherwise.

- 1) Uncertainty about true distribution (24 Points): Consider a source U with alphabet $U = \{a_1, \ldots, a_m\}$ and suppose we know that the true distribution of U is either P_1 or P_2 , but we are not sure which.
 - a) (8 points) **True/False:** There is a prefix code where the length of the codeword associated to a_i is $l_i = \left\lceil \log_2 \left(\frac{2}{P_1(a_i) + P_2(a_i)} \right) \right\rceil$. Solution:

Let $l_i = \left\lceil \log_2 \left(\frac{2}{P_1(a_i) + P_2(a_i)} \right) \right\rceil$, and compute the Kraft sum:

$$\sum_{m=1}^{M} 2^{-l_i} \le \sum_{m=1}^{M} 2^{-\log_2\left(\frac{2}{P_1(a_i) + P_2(a_i)}\right)}$$
$$= \sum_{m=1}^{M} \frac{P_1(a_i) + P_2(a_i)}{2}$$
$$= 1.$$

Accordingly, the Kraft sum at most to 1, and therefore, there exists a prefix-free code where the length of the codeword associated to a_i is l_i .

b) (8 points) Show that the average (computed using the true distribution) length \bar{l} of the code constructed in item (a) satisfies $H(U) \leq \bar{l} \leq H(U) + 2$.

Solution:

Since the constructed code in item (a) is a prefix code, then $l \ge H(U)$. To prove the upper bound, let P^* be the true distribution (which is either P_1 or P_2). Then, clearly, the inequality $P^*(a_i) \le P_1(a_i) + P_2(a_i)$ holds for all $1 \le i \le M$. Accordingly, we get:

$$\begin{split} \bar{l} &= \sum_{m=1}^{M} P^*(a_i) l_i \\ &= \sum_{m=1}^{M} P^*(a_i) \left\lceil \log_2 \left(\frac{2}{P_1(a_i) + P_2(a_i)}\right) \right\rceil \\ &= \sum_{m=1}^{M} P^*(a_i) \left(1 + \log_2 \left(\frac{2}{P_1(a_i) + P_2(a_i)}\right)\right) \\ &= 2 + \sum_{m=1}^{M} P^*(a_i) \log_2 \left(\frac{1}{P_1(a_i) + P_2(a_i)}\right) \\ &\leq 2 + \sum_{m=1}^{M} P^*(a_i) \log_2 \left(\frac{1}{P^*(a_i)}\right) \\ &= 2 + H(U). \end{split}$$

c) (8 points) Now assume that the true distribution of U is one of k distributions P_1, \ldots, P_k , but we don't know which. Show that there exists a prefix code satisfying $H(U) \le \overline{l} \le H(U) + \log_2(k) + 1$. Solution:

Now let $l_i = \left\lceil \log_2 \left(\frac{k}{P_1(a_i) + \dots + P_k(a_i)} \right) \right\rceil$, and let us compute the Kraft sum for this scenario:

$$\sum_{m=1}^{M} 2^{-l_i} \le \sum_{m=1}^{M} 2^{-\log_2\left(\frac{k}{P_1(a_i) + \dots + P_k(a_i)}\right)}$$
$$= \sum_{m=1}^{M} \frac{P_1(a_i) + \dots + P_k(a_i)}{k}$$
$$= 1.$$

Thus, the code is a prefix code which implies that $\overline{l} \ge H(U)$. Here too, let P^* denote the true distribution. Therefore, $P^*(a_i) \le P_1(a_i) + \cdots + P_k(a_i)$ for all $1 \le i \le M$. Following the same proof steps as for the previous item, it is easy to show that:

$$l \le 1 + \log_2(k) + H(U).$$

2) **GMM (18 points):** We will derive the EM update rules for a univariate Gaussian Mixture Model with two mixture components. The mean μ will be shared between the two mixture components, but each component will have its own standard deviation σ_k . The model will be defined as follows:

$$z \sim Bernoulli(\theta),$$

 $p(x|z=k) \text{ is } \mathcal{N}(\mu, \sigma_k).$

a) (4 points) Write the density defined by this model (i.e. the probability of x, with z marginalized out)

Solution:

$$p(x) = \theta \mathcal{N}(x; \mu, \sigma_1) + (1 - \theta) \mathcal{N}(x; \mu, \sigma_0)$$

b) (4 points) E-step - Compute the posterior probability $w^{(i)} = Pr(z^{(i)} = 1|x^{(i)})$

Solution:

$$w^{(i)} = \frac{\theta \mathcal{N}(x; \mu, \sigma_1)}{\theta \mathcal{N}(x; \mu, \sigma_1) + (1 - \theta) \mathcal{N}(x; \mu, \sigma_0)}$$

- c) (5 points) M-Step Calculate the update rule for μ (for a fixed σ_k)
- d) (5 points) M-Step Calculate the update rule for σ_k (for a fixed μ)

Solution:

At each M-step we optimize the following:

$$\begin{split} \mathcal{L}(\mu, \sigma_0, \sigma_1, \theta) &= \sum_{i=1}^N w^{(i)} \log \left(\mathcal{N}(x(i)|\mu, \sigma_1) \right) + w^{(i)} \log \theta \\ &+ (1 - w^{(i)}) \log \left(\mathcal{N}(x(i)|\mu, \sigma_0) \right) + (1 - w^{(i)}) \log(1 - \theta) \\ \frac{\partial \mathcal{L}}{\partial \mu} &= 0 \quad \Rightarrow \quad \sum_i^N (w^{(i)} \frac{x^{(i)} - \mu}{\sigma_1^2} + (1 - w^{(i)}) \frac{x^{(i)} - \mu}{\sigma_0^2} = 0 \\ &\Rightarrow \quad \sum_i^N (x^{(i)} - \mu) \left(\frac{w^{(i)}}{\sigma_1^2} + \frac{1 - w^{(i)}}{\sigma_0^2} \right) = 0 \\ &\Rightarrow \quad \sum_i^N (x^{(i)} - \mu) \left(\sigma_0^2 w^{(i)} + \sigma_1^2 (1 - w^{(i)}) \right) = 0 \end{split}$$

Thus you get:

$$\mu = \frac{\sum_{i}^{N} x^{(i)} \left(\sigma_{0}^{2} w^{(i)} + \sigma_{1}^{2} (1 - w^{(i)})\right)}{\sum_{i}^{N} \left(\sigma_{0}^{2} w^{(i)} + \sigma_{1}^{2} (1 - w^{(i)})\right)}$$

$$\frac{\partial \mathcal{L}}{\sigma_{k}^{2}} = 0 \Rightarrow \sigma_{k}^{2} = \frac{\sum_{i=1}^{N} w^{(i)} (x^{(i)} - \mu)^{2}}{\sum_{i=1}^{N} w^{(i)}}$$

3) Linear Regression (26 Points): You are tasked with solving a fitting a linear regression model on a set of m datapoints where each feature has some dimensionality d. Your dataset can be described as the set {x⁽ⁱ⁾, y⁽ⁱ⁾}_{i=1}^m, where x⁽ⁱ⁾ ∈ ℝ^d, y⁽ⁱ⁾ ∈ ℝ. You initially decide to optimize the loss objective:

$$J = \frac{1}{m} \sum_{i=1}^{m} (y^{(i)} - x^{(i)^{T}} \theta)^{2},$$

using Batch Gradient Descent - in which each step involves calculations over the **entire** training set. Here, $\theta \in \mathbb{R}^d$ is your weight vector. Assume you are ignoring a bias term for this problem.

a) (4 points) Write each update of the batch gradient descent, $\frac{\partial J}{\partial \theta}$ in **vectorized** form. Your solution should be a single vector (no summation terms) in terms of the matrix X and vectors Y and θ , where

$$X = \begin{bmatrix} x^{(1)^T} \\ \vdots \\ x^{(m)^T} \end{bmatrix}, Y = \begin{bmatrix} y^{(1)} \\ \vdots \\ y^{(m)} \end{bmatrix}.$$

Solution:

Final solution is:

$$\frac{\partial J}{\partial \theta} = \frac{2}{m} X^T (X\theta - Y)$$

Two common approaches are "derive, then vectorize" and "vectorize, then derive". Both get full credit. With the first approach.

$$\frac{\partial J}{\partial \theta} = \frac{2}{m} \sum_{i} (x^{(i)^{T}} \theta - y^{(i)}) x^{(i)} - \text{derivative step}$$
$$= \frac{\partial J}{\partial \theta} = \frac{2}{m} X^{T} (X \theta - Y) - \text{vectorization step}$$

b) (7 points) A coworker suggests you augment your dataset by adding Gaussian noise to your features. Specifically, you would be adding *zero-mean*, Gaussian noise of *known vairance* σ^2 from the distribution

$$\mathcal{N}(0,\sigma^2 I),$$

where $I \in \mathbb{R}^{d \times d}$, $\sigma \in \mathbb{R}$. This modifies your original objective to:

$$J_* = \frac{1}{m} \sum_{i=1}^m (y^{(i)} - (x^{(i)} - \delta^{(i)})^T \theta)^2,$$

where $\delta^{(i)}$ are **i.i.d.** noise vectors, $\delta^{(i)} \in \mathbb{R}^d$ and $\delta^{(i)} \sim \mathcal{N}(0, \sigma^2 I)$.

Express the expectation of the modified objective J_* over the Gaussian noise, $\mathbb{E}_{\delta \sim \mathcal{N}}[J_*]$, as a function of the original objective J added to a term independent of your data. Your answer should be in the form

$$\mathbb{E}_{\delta \sim \mathcal{N}}[J_*] = J + C$$

where C is independent of points in $\{x^{(i)}, y^{(i)}\}_{i=1}^{m}$. **Hint:** For a Gaussian random vector δ with zero mean, and convariance matrix $\sigma^2 I$

$$\mathbb{E}_{\delta \sim \mathcal{N}}[\delta \delta^T] = \sigma^2 I, \quad \mathbb{E}_{\delta \sim \mathcal{N}}[\delta] = 0.$$

Solution:

$$J_* = \frac{1}{m} \sum_{i=1}^m (y^{(i)} - (x^{(i)} - \delta^{(i)})^T \theta)^2$$

= $\frac{1}{m} \sum_{i=1}^m ((y^{(i)} - x^{(i)}) - \delta^{(i)})^T \theta)^2$
= $\frac{1}{m} \sum_{i=1}^m ((y^{(i)} - x^{(i)})^2 - 2(y^{(i)} - x^{(i)})(\delta^{(i)T}\theta) + (\delta^{(i)T}\theta)^2))$
= $J + \frac{1}{m} \sum_{i=1}^m (-2(y^{(i)} - x^{(i)})(\delta^{(i)T}\theta) + (\delta^{(i)T}\theta)^2)$

$$\mathbb{E}_{\delta \sim \mathcal{N}}[J_*] = J + \mathbb{E}_{\delta \sim \mathcal{N}} \left[\frac{1}{m} \sum_{i=1}^m (-2(y^{(i)} - x^{(i)})(\delta^{(i)T}\theta) + (\delta^{(i)T}\theta)^2) \right]$$

From Linearity of Expectation, we can take the expectation individually for each sample:

$$\mathbb{E}_{\delta \sim \mathcal{N}} \left[\frac{1}{m} \sum_{i=1}^{m} -2(y^{(i)} - x^{(i)})(\delta^{(i)T}\theta) \right] = -2(y^{(i)} - x^{(i)})\mathbb{E} \left[\delta^{(i)T}\theta \right] = 0$$

and
$$\mathbb{E}_{\delta \sim \mathcal{N}} \left| (\delta^{(i)T} \theta)^2 \right| = \sigma^2 \|\theta\|_2^2$$
 (from the hint)

Thus, we get:

$$\mathbb{E}_{\delta \sim \mathcal{N}}[J_*] = J + \sigma^2 \|\theta\|_2^2. \tag{1}$$

4

c) (4 points) What effect would adding noise have on model overfitting/underfitting? Explain why.

Remember that the weights update rule is derived from the loss function, which is the expectation of J_* .

Solution:

Adding noise to the model will prevent overfiting because the model wouldn't be able to remember a specific point mapping from $x^{(i)}$ to $y^{(i)}$ due to the noise inserted to $x^{(i)}$. Alternatively, in expectation, the new objective would help regularize the model, due to the similar L2 regularization term (see the answer for next question) and it is known that L2 regularization method prevents overfiting.

d) (4 points) Is this method similar to a regularization method we studied in class? If so, specify the regularization method and prove it and if not, explain why?

Solution:

Yes. It is similar to L_2 regularization, but with a scalar multiplicative σ^2 (see equation 1) which is the λ in L2 regularization method.

e) (3 points) Consider the limits $\sigma \to 0$ and $\sigma \to \infty$. What impact would these extremes in the value of σ have on model training (relative to no noise added)? Explain why.

Solution:

- $\sigma \rightarrow 0:$ Less regularizing/no effect.
- $\sigma \rightarrow \infty:$ All weights get pushed to zero / model underfits.
- f) (4 points) Suggest a cost function and a noise that is related to *Dropout*. **Solution:**

$$J_* = \frac{1}{m} \sum_{i=1}^m (y^{(i)} - (x^{(i)} \cdot \gamma^{(i)})^T \theta)^2$$

where $\gamma^{(i)} \in \mathbb{R}^d$ are i.i.d. noise vectors. $\gamma_j^{(i)} \sim Ber(p)$ (i.i.d.) for $j \in \{1, 2, ..., d\}$, and the probability 1 - p is the rate of the dropout.

4) Computable lower bounds (32 Points): In this question, you will prove a simple lower bound on the capacity of a memoryless channel. Let p(y|x) be a memoryless channel, and let p(x) be a distribution on X. Let r(x|y) be an arbitrary conditional distribution on X given Y, i.e., for each x ∈ X and each y ∈ Y, r(x|y) ≥ 0 and ∑_{x∈X} r(x|y) = 1. Define the functional F(p, r) as follows:

$$F(p,r) = \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} p(x)p(y|x) \log_2\left(\frac{r(x|y)}{p(x)}\right).$$

where p in F(p,r) denotes P(x) and p(y|x) is fixed thought the question. Now, for each input distribution p on \mathcal{X} , define the conditional distribution r_p as

$$r_p(x|y) = \frac{p(x)p(y|x)}{\sum_{\tilde{x} \in \mathcal{X}} p(\tilde{x})p(y|\tilde{x})}$$

That is, r_p is the "true" conditional distribution of \mathcal{X} given \mathcal{Y} when p is the input distribution.

a) (8 points) **True/False:** For all conditional distributions r we have $F(p, r) \le F(p, r_p)$. Solution:

True. Let us show that the difference is non-negative.

$$F(p, r_p) - F(p, r) = \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} p(x) p(y|x) \log_2 \left(\frac{r_p(x|y)}{r(x|y)}\right)$$
$$= \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} p(x) p(y|x) \log_2 \left(\frac{p(x) p(y|x)}{r(x|y) \sum_{\tilde{x} \in \mathcal{X}} p(\tilde{x}) p(y|\tilde{x})}\right)$$
$$= D(P_1 ||P_2)$$
$$\ge 0$$

where $P_1(x,y) = p(x)p(y|x)$ and $P_2(x,y) = r(x|y)\sum_{\tilde{x}\in\mathcal{X}} p(\tilde{x})p(y|\tilde{x})$.

b) (4 points) Show that $I(X;Y) = \max_{r} F(p,r)$. Solution: From the previous item we can deduce that $F(p, r_p) = \max_r F(p, r)$. Further,

$$F(p, r_p) = \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} p(x)p(y|x) \log_2\left(\frac{r_p(x|y)}{p(x)}\right)$$
$$= \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} p(x)p(y|x) \log_2\left(\frac{p(x,y)}{p(x)p(y)}\right)$$
$$= I(X; Y),$$

as required.

c) (8 points) **True/False:** The functional F(p, r) is strictly concave in both p and r. Solution:

True. We can rewrite F(p, r) as follows:

$$F(p,r) = \left(\sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} p(x)p(y|x)\log_2(r(x|y))\right) + \left(\sum_{x \in \mathcal{X}} p(x)\log_2\left(\frac{1}{p(x)}\right)\right)$$

The first term is linear in p while the second term is strictly concave in p (the function $t \Rightarrow t \log_2 \frac{1}{t}$ is strictly concave). Therefore, F(p,r) is strictly concave in p. In addition, the first term is concave in r (the function \log_2 is strictly concave), and the second term is constant with respect to r. Therefore, F(p,r) is strictly concave in r.

d) (6 points) In Algorithm 1 below, we introduce an iterative algorithm for maximizing a two-variable function. Following the previous items, suggest such an iterative algorithm to compute the capacity.

Solution:

Following the previous items, the capacity can be computed as

$$C = \max_{p} I(X;Y) \tag{2}$$

$$= \max_{p} \max_{r} F(p, r).$$
(3)

Accordingly, in the spirit of Algorithm 1, the following algorithm can be used to compute the capacity:

Algorithm 1 Alternating maximization procedure

input: The function F(p, r) that is concave in both p and r **output:** A global maximum of F(p, r) (capacity) set p_0 uniform in \mathcal{X} and solve $r_0 = \arg \max_r F(p_0, r)$ set i = 1while $F(p_i, r_i)$ not converged do $p_i = \arg\max_p F(p, r_{i-1})$ $r_i = \arg \max_r F(p_{i-1}, r)$ compute $F(p_i, r_i)$ i = i + 1end return $F(p_i, r_i)$

> **Note:** The Alternating maximization procedure is known to converge to optimal solution when the function g(x, y) is concave in (x, y).

e) (6 points) For a given memoryless channel, let r^* denote the conditional distribution that should be used to obtain the capacity. Write explicitly r^* for the case of a binary symmetric channel with crossover probability 0.2. Solution:

The optimal input distribution p^* for a binary symmetric channel is a uniform distribution. Accordingly,

$$r^*(x|y) = r_{p^*}(x|y)$$
$$= \frac{p^*(x)p(y|x)}{\sum_{\tilde{x}\in\mathcal{X}} p^*(\tilde{x})p(y|\tilde{x})}.$$

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Good Luck!