Solutions to Set #2
Data Compression, Huffman code and AEP

1. Huffman coding.
Consider the random variable
\[ X = \left( \begin{array}{cccccc}
 x_1 & x_2 & x_3 & x_4 & x_5 & x_6 & x_7 \\
 0.50 & 0.26 & 0.11 & 0.04 & 0.04 & 0.03 & 0.02 \\
\end{array} \right) \]

(a) Find a binary Huffman code for \( X \).
(b) Find the expected code length for this encoding.
(c) Extend the Binary Huffman method to Ternary (Alphabet of 3) and apply it for \( X \).

Solution: Huffman coding.

(a) The Huffman tree for this distribution is

```
Codeword
1  x_1  0.50  0.50  0.50  0.50  0.50  1
01  x_2  0.26  0.26  0.26  0.26  0.26  0.50
001  x_3  0.11  0.11  0.11  0.11  0.24
00011  x_4  0.04  0.04  0.08  0.13
00010  x_5  0.04  0.04  0.05
00001  x_6  0.03  0.05
00000  x_7  0.02
```

(b) The expected length of the codewords for the binary Huffman code is 2 bits. \( H(X) = 1.99 \) bits

(c) The ternary Huffman tree is

```
Codeword
0  x_1  0.50  0.50  0.50  1.0
1  x_2  0.26  0.26  0.26
20  x_3  0.11  0.11  0.24
21  x_4  0.04  0.04
222  x_5  0.04  0.09
221  x_6  0.03
220  x_7  0.02
```
This code has an expected length 1.33 ternary symbols. \( H_3(X) = 1.25 \) ternary symbols.

1
2. **Codes.**

Let \( X_1, X_2, \ldots \), i.i.d. with

\[
X = \begin{cases} 
1, & \text{with probability } \frac{1}{2} \\
2, & \text{with probability } \frac{1}{4} \\
3, & \text{with probability } \frac{1}{4}.
\end{cases}
\]

Consider the code assignment

\[
C(x) = \begin{cases} 
0, & \text{if } x = 1 \\
01, & \text{if } x = 2 \\
11, & \text{if } x = 3.
\end{cases}
\]

(a) Is this code nonsingular?
(b) Uniquely decodable?
(c) Instantaneous?
(d) Entropy Rate is defined as

\[
H(X) \triangleq \lim_{n \to \infty} \frac{H(X^n)}{n}.
\]

What is the entropy rate of the process

\( Z_1Z_2Z_3 \ldots = C(X_1)C(X_2)C(X_3) \ldots \)?

**Solution: Codes.**

(a) **Yes**, this code is nonsingular because \( C(x) \) is different for every \( x \).

(b) **Yes**, this code is uniquely decodable. Reversing the codewords

\[
C'(x) = \begin{cases} 
0, & \text{if } x = 1 \\
10, & \text{if } x = 2 \\
11, & \text{if } x = 3
\end{cases}
\]

gives an instantaneous code, and thus a uniquely decodable code. Therefore the reversed extension is uniquely decodable, and so the extension itself is also uniquely decodable.
(c) **No**, this code is not instantaneous because \( C(1) \) is a prefix of \( C(2) \).

(d) The expected codeword length is

\[
L(C(x)) = 0.5 \times 1 + 0.25 \times 2 + 0.25 \times 2 = \frac{3}{2}.
\]

Further, the entropy rate of the i.i.d. \( X^n \) is

\[
H(X) = H(X) = H(0.5, 0.25, 0.25) = \frac{3}{2}.
\]

So the code is a uniquely decodable code with \( L = H(X) \), and therefore the sequence is maximally compressed with \( H(Z) = 1 \) bit. If \( H(Z) \) were less than its maximum of 1 bit then the \( Z^n \) sequence could be further compressed to its entropy rate, and \( X^n \) could also be compressed further by blockcoding. However, this would result in \( L_m < H(X) \) which contradicts theorem 5.4.2 of the text. So \( H(Z) = 1 \) bit.

Note that the \( Z^n \) sequence is not i.i.d. \( \sim \text{Br}(\frac{1}{2}) \), even though \( H(Z) = 1 \) bit. For example, \( P\{Z_1 = 1\} = \frac{1}{7} \), and a sequence starting 10... is not allowed. However, once \( Z_i = 0 \) for some \( i \) then \( Z_k \) is Bernoulli(\( \frac{1}{2} \)) for \( k > i \), so \( Z^n \) is asymptotically Bernoulli(\( \frac{1}{2} \)) and gives the entropy rate of 1 bit.

3. Huffman via MATLAB

(a) Give a Huffman encoding into an alphabet of size \( D = 2 \) of the following probability mass function:

\[
\left( \frac{1}{2}, \frac{1}{8}, \frac{1}{8}, \frac{1}{16}, \frac{1}{16}, \frac{1}{16}, \frac{1}{16} \right)
\]

(b) Assume you have a file of size 1,000 symbols where the symbols are distributed i.i.d. according to the pmf above. After applying the Huffman code, what would be the pmf of the compressed binary file and what would be the expected length?

(c) Generate a sequence (using MATLAB or any other software) of length 10,000 symbols of \( X \) with i.i.d probability \( P_X \). Assume the alphabet of \( X \) is \( \mathcal{X} = \{0, 1, \ldots, 6\} \).
(d) What is the percentage of each symbol (0, 1, . . . , 6) in the sequence that was generated by MATLAB. Explain this result using the law of large numbers.

(e) Represent each symbol in $\mathcal{X}$ using the simple binary representation. Namely, $X = 0$ represent as '000', $X = 1$ represent as '001', $X = 2$ represent as '010', . . . , $X = 6$ represent as '110'.

(f) What is the length of the simple representation. What percentage of '0' and '1' do you have in this representation?

(g) Now, compress the 10,000 symbols of $X$, into bits using Huffman code.

(h) What is the length of the compressed file? What is the percentage of '0' and '1' do you have in this representation?

(i) Explain the results.

Solution:

(a) The code is presented in Fig 1.

(b) Figure 1: Huffman
(b) Huffman code is optimal code and achieves the entropy for dyadic distribution. If the distribution of the digits is not $Bernoulli(\frac{1}{2})$ you can compress it further. The binary digits of the data would be equally distributed after applying the Huffman code and therefore $p_0 = p_1 = \frac{1}{2}$.

The expected length would be:

$$E[l] = \frac{1}{2} \cdot 1 + \frac{1}{8} \cdot 3 + \frac{1}{8} \cdot 3 + \frac{1}{16} \cdot 4 + \frac{1}{16} \cdot 4 + \frac{1}{16} \cdot 4 + \frac{1}{16} \cdot 4 = 2.25$$

Therefore, the expected length of 1000 symbols would be 2250 bits.

(c-i) The results from simulation should be as follows:

<table>
<thead>
<tr>
<th>Drawn symbols</th>
<th>Size: 10000</th>
</tr>
</thead>
<tbody>
<tr>
<td>Percentage of x=0:</td>
<td>0.4984</td>
</tr>
<tr>
<td>Percentage of x=1:</td>
<td>0.125</td>
</tr>
<tr>
<td>Percentage of x=2:</td>
<td>0.1262</td>
</tr>
<tr>
<td>Percentage of x=3:</td>
<td>0.0636</td>
</tr>
<tr>
<td>Percentage of x=4:</td>
<td>0.0575</td>
</tr>
<tr>
<td>Percentage of x=5:</td>
<td>0.0647</td>
</tr>
<tr>
<td>Percentage of x=6:</td>
<td>0.0646</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Before compression</th>
<th>Size: 30000</th>
</tr>
</thead>
<tbody>
<tr>
<td>Percentage of b=0:</td>
<td>0.77213</td>
</tr>
<tr>
<td>Percentage of b=1:</td>
<td>0.22787</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>After compression</th>
<th>Size: 22463</th>
</tr>
</thead>
<tbody>
<tr>
<td>Percentage of b=0:</td>
<td>0.49958</td>
</tr>
<tr>
<td>Percentage of b=1:</td>
<td>0.50042</td>
</tr>
</tbody>
</table>

Compressions rate (symbol -> binary): 2.2463
Explanation: The symbols $X(i)$, $i \in \{1, 2, \ldots, 10,000\}$ are drawn i.i.d. according to $P_X$. Therefore, by the L.L.N, the percentage of appearances of each symbol $x \in X$, should be approximately $P_X(x)$. Next, we represent each symbol with 3 binary symbols. Since the file in uncompressed, the percentage of 0’s is not half. The expected percentage can be calculated from the PMF. In our case,

$$
\mathbb{E}[\#\{0\text{’s in file}\}] = 10,000(P_x(0) \cdot 3 + P_x(1) \cdot 2 + P_x(2) \cdot 2 + P_x(3) \cdot 1 + \cdots + P_x(6) \cdot 1)
$$

$$
= 23,125
$$

Since there are 30,000 bits for representation, the expected percentage is 77.08%. After Huffman’s code is applied, the number of appearances of 0’s and 1’s are almost equal. This is since lossless compression maximize entropy (now 0 and 1 in the file are uniform).

4. **Entropy and source coding of a source with infinite alphabet**
    
    (15 points)

    Let $X$ be an i.i.d. random variable with an infinite alphabet, $X = \{1, 2, 3, \ldots\}$. In addition let $P(X = i) = 2^{-i}$.

    (a) What is the entropy of the random variable?

    (b) Find an optimal variable length code, and show that it is indeed optimal.

**Solution**

(a) 

$$
H(X) = -\sum_{x \in X} p(x) \log p(x)
$$

$$
= -\sum_{i=1}^{\infty} 2^{-i} \log_2(2^{-i})
$$

$$
= -\sum_{i=1}^{\infty} \frac{-i}{2^i} = 2
$$
(b) Coding Scheme:
\begin{align*}
1 & \quad 0 \\
2 & \quad 10 \\
3 & \quad 110 \\
4 & \quad 1110 \\
5 & \quad 11110 \\
\vdots & \quad \vdots \\
\vdots & \quad \vdots \\
\end{align*}

Average Length:
\[ L^* = \sum_{i=1}^{\infty} p(x = i)L(i) = \sum_{i=1}^{\infty} \frac{i}{2^i} = 2 = H(X) \]

Hence it is the Optimal Code.

5. **Bad wine.**

One is given 6 bottles of wine. It is known that precisely one bottle has gone bad (tastes terrible). From inspection of the bottles it is determined that the probability \( p_i \) that the \( i \)th bottle is bad is given by \((p_1, p_2, \ldots, p_6) = (\frac{7}{26}, \frac{5}{26}, \frac{4}{26}, \frac{4}{26}, \frac{3}{26}, \frac{3}{26})\). Tasting will determine the bad wine.

Suppose you taste the wines one at a time. Choose the order of tasting to minimize the expected number of tastings required to determine the bad bottle. Remember, if the first 5 wines pass the test you don’t have to taste the last.

(a) What is the expected number of tastings required?
(b) Which bottle should be tasted first?

Now you get smart. For the first sample, you mix some of the wines in a fresh glass and sample the mixture. You proceed, mixing and tasting, stopping when the bad bottle has been determined.

(c) What is the minimum expected number of tastings required to determine the bad wine?
(d) What mixture should be tasted first?
Solution: Bad wine.

(a) If we taste one bottle at a time, the corresponding number of tastings are \(\{1, 2, 3, 4, 5, 5\}\) with some order. By the same argument as in Lemma 5.8.1, to minimize the expected length \(\sum p_i l_i\) we should have \(l_j \leq l_k\) if \(p_j > p_k\). Hence, the best order of tasting should be from the most likely wine to be bad to the least.

The expected number of tastings required is

\[
\sum_{i=1}^{6} p_i l_i = 1 \times \frac{7}{26} + 2 \times \frac{5}{26} + 3 \times \frac{4}{26} + 4 \times \frac{4}{26} + 5 \times \frac{3}{26} + 5 \times \frac{3}{26}
\]

\[
= \frac{75}{26}
\]

\[
= 2.88
\]

(b) The first bottle to be tasted should be the one with probability \(\frac{7}{26}\).

(c) The idea is to use Huffman coding.

\[
(01) \quad 7 \quad 7 \quad 8 \quad 11 \quad 15 \quad 26 \\
(11) \quad 5 \quad 6 \quad 7 \quad 8 \quad 11 \\
(000) \quad 4 \quad 5 \quad 6 \quad 7 \\
(001) \quad 4 \quad 4 \quad 5 \\
(100) \quad 3 \quad 4 \\
(101) \quad 3
\]

The expected number of tastings required is

\[
\sum_{i=1}^{6} p_i l_i = 2 \times \frac{7}{26} + 2 \times \frac{5}{26} + 3 \times \frac{4}{26} + 3 \times \frac{4}{26} + 3 \times \frac{3}{26} + 3 \times \frac{3}{26}
\]

\[
= \frac{66}{26}
\]

\[
= 2.54
\]

Note that \(H(p) = 2.52\) bits.

(d) The mixture of the first, third, and forth bottles should be tasted first, (or equivalently the mixture of the second, fifth and sixth).
6. Relative entropy is cost of miscoding.
Let the random variable $X$ have five possible outcomes $\{1, 2, 3, 4, 5\}$. Consider two distributions on this random variable

<table>
<thead>
<tr>
<th>Symbol</th>
<th>$p(x)$</th>
<th>$q(x)$</th>
<th>$C_1(x)$</th>
<th>$C_2(x)$</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>$\frac{1}{2}$</td>
<td>$\frac{1}{2}$</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>2</td>
<td>$\frac{1}{4}$</td>
<td>$\frac{1}{8}$</td>
<td>10</td>
<td>100</td>
</tr>
<tr>
<td>3</td>
<td>$\frac{1}{8}$</td>
<td>$\frac{1}{8}$</td>
<td>110</td>
<td>101</td>
</tr>
<tr>
<td>4</td>
<td>$\frac{1}{16}$</td>
<td>$\frac{1}{8}$</td>
<td>1110</td>
<td>110</td>
</tr>
<tr>
<td>5</td>
<td>$\frac{1}{16}$</td>
<td>$\frac{1}{8}$</td>
<td>1111</td>
<td>111</td>
</tr>
</tbody>
</table>

(a) Calculate $H(p)$, $H(q)$, $D(p||q)$ and $D(q||p)$.

(b) The last two columns above represent codes for the random variable. Verify that the average length of $C_1$ under $p$ is equal to the entropy $H(p)$. Thus $C_1$ is optimal for $p$. Verify that $C_2$ is optimal for $q$.

(c) Now assume that we use code $C_2$ when the distribution is $p$. What is the average length of the codewords. By how much does it exceed the entropy $H(p)$?

(d) What is the loss if we use code $C_1$ when the distribution is $q$?

Solution: Relative entropy is cost of miscoding.

(a)

$$H(p) = \sum_i -p_i \log p_i$$

$$= -\frac{1}{2} \log \frac{1}{2} - \frac{1}{4} \log \frac{1}{4} - \frac{1}{8} \log \frac{1}{8} - 2 \cdot \frac{1}{16} \log \frac{1}{16}$$

$$= \frac{15}{8}.$$  

Similarly, $H(q) = 2$.

$$D(p||q) = \sum_i p_i \log \frac{p_i}{q_i}$$

$$= \frac{1}{2} \log \frac{1}{1/2} + \frac{1}{4} \log \frac{1/4}{1/8} + \frac{1}{8} \log \frac{1/8}{1/8} + 2 \cdot \frac{1}{16} \log \frac{1/16}{1/8}$$

$$= \frac{1}{8}.$$
Similarly, \( D(q||p) = \frac{1}{8} \).

(b) The average codeword length for \( C_1 \) is

\[
El_1 = \frac{1}{2} \cdot 1 + \frac{1}{4} \cdot 2 + \frac{1}{8} \cdot 3 + \frac{1}{16} \cdot 4 = \frac{15}{8}.
\]

Similarly, the average codeword length for \( C_2 \) is 2.

(c) \( E_{p}l_2 = \frac{1}{2} \cdot 1 + \frac{1}{4} \cdot 3 + \frac{1}{8} \cdot 3 + \frac{1}{16} \cdot 3 = 2 \),

which exceeds \( H(p) \) by \( D(p||q) = \frac{1}{8} \).

(d) Similarly, \( E_{q}l_1 = \frac{17}{8} \), which exceeds \( H(q) \) by \( D(q||p) = \frac{1}{8} \).

7. Shannon code. Consider the following method for generating a code for a random variable \( X \) which takes on \( m \) values \( \{1, 2, \ldots, m\} \) with probabilities \( p_1, p_2, \ldots, p_m \). Assume that the probabilities are ordered so that \( p_1 \geq p_2 \geq \cdots \geq p_m \). Define

\[
F_i = \sum_{k=1}^{i-1} p_i,
\]

the sum of the probabilities of all symbols less than \( i \). Then the codeword for \( i \) is the number \( F_i \in [0, 1] \) rounded off to \( l_i \) bits, where \( l_i = \lceil \log \frac{1}{p_i} \rceil \).

(a) Show that the code constructed by this process is prefix-free and the average length satisfies

\[
H(X) \leq L < H(X) + 1.
\]

(b) Construct the code for the probability distribution \((0.5, 0.25, 0.125, 0.125)\).

Solution to Shannon code.

(a) Since \( l_i = \lceil \log \frac{1}{p_i} \rceil \), we have

\[
\log \frac{1}{p_i} \leq l_i < \log \frac{1}{p_i} + 1
\]
which implies that

\[ H(X) \leq L = \sum p_i l_i < H(X) + 1. \]  

(5)

The difficult part is to prove that the code is a prefix code. By the choice of \( l_i \), we have

\[ 2^{-l_i} \leq p_i < 2^{-(l_i - 1)}. \]  

(6)

Thus \( F_j, j > i \) differs from \( F_i \) by at least \( 2^{-l_i} \), and will therefore differ from \( F_i \) is at least one place in the first \( l_i \) bits of the binary expansion of \( F_i \). Thus the codeword for \( F_j, j > i \), which has length \( l_j \geq l_i \), differs from the codeword for \( F_i \) at least once in the first \( l_i \) places. Thus no codeword is a prefix of any other codeword.

(b) We build the following table

<table>
<thead>
<tr>
<th>Symbol</th>
<th>Probability</th>
<th>( F_i ) in decimal</th>
<th>( F_i ) in binary</th>
<th>( l_i )</th>
<th>Codeword</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>0.5</td>
<td>0.0</td>
<td>0.0</td>
<td>1</td>
<td>0</td>
</tr>
<tr>
<td>2</td>
<td>0.25</td>
<td>0.5</td>
<td>0.10</td>
<td>2</td>
<td>10</td>
</tr>
<tr>
<td>3</td>
<td>0.125</td>
<td>0.75</td>
<td>0.110</td>
<td>3</td>
<td>110</td>
</tr>
<tr>
<td>4</td>
<td>0.125</td>
<td>0.875</td>
<td>0.111</td>
<td>3</td>
<td>111</td>
</tr>
</tbody>
</table>

The Shannon code in this case achieves the entropy bound (1.75 bits) and is optimal.

8. **An AEP-like limit.** Let \( X_1, X_2, \ldots \) be i.i.d. drawn according to probability mass function \( p(x) \). Find

\[ \lim_{n \to \infty} [p(X_1, X_2, \ldots, X_n)]^{\frac{1}{n}}. \]

**Solution:** An AEP-like limit.

\( X_1, X_2, \ldots \), i.i.d. \( \sim p(x) \). Hence \( \log(X_i) \) are also i.i.d. and

\[ \lim(p(X_1, X_2, \ldots, X_n))^{\frac{1}{n}} = \lim 2^{\log(p(X_1, X_2, \ldots, X_n))^{\frac{1}{n}}} \]

\[ = 2^{\lim \frac{1}{n} \sum \log p(X_i)} \]

\[ = 2^{E(\log(p(X)))} \]

\[ = 2^{-H(X)} \]

by the strong law of large numbers.
9. **AEP.** Let $X_1, X_2, \ldots$ be independent identically distributed random variables drawn according to the probability mass function $p(x), x \in \{1, 2, \ldots, m\}$. Thus $p(x_1, x_2, \ldots, x_n) = \prod_{i=1}^n p(x_i)$. We know that $-\frac{1}{n} \log p(X_1, X_2, \ldots, X_n) \to H(X)$ in probability. Let $q(x_1, x_2, \ldots, x_n) = \prod_{i=1}^n q(x_i)$, where $q$ is another probability mass function on $\{1, 2, \ldots, m\}$.

(a) Evaluate $\lim_{n \to \infty} -\frac{1}{n} \log q(X_1, X_2, \ldots, X_n)$, where $X_1, X_2, \ldots$ are i.i.d. $\sim p(x)$.

(b) Now evaluate the limit of the log likelihood ratio $\frac{1}{n} \log \frac{q(X_1, \ldots, X_n)}{p(X_1, \ldots, X_n)}$ when $X_1, X_2, \ldots$ are i.i.d. $\sim p(x)$. Thus the odds favouring $q$ are exponentially small when $p$ is true.

**Solution: AEP.**

(a) Since the $X_1, X_2, \ldots, X_n$ are i.i.d., so are $q(X_1), q(X_2), \ldots, q(X_n)$, and hence we can apply the strong law of large numbers to obtain

$$\lim_{n \to \infty} -\frac{1}{n} \log q(X_1, X_2, \ldots, X_n) = \lim_{n \to \infty} -\frac{1}{n} \sum \log q(X_i)$$

$$= -E(\log q(X)) \text{ w.p. 1}$$

$$= - \sum p(x) \log q(x)$$

$$= \sum p(x) \log \frac{p(x)}{q(x)} - \sum p(x) \log p(x)$$

$$= D(p||q) + H(p).$$

(b) Again, by the strong law of large numbers,

$$\lim_{n \to \infty} -\frac{1}{n} \log \frac{q(X_1, X_2, \ldots, X_n)}{p(X_1, X_2, \ldots, X_n)} = \lim_{n \to \infty} -\frac{1}{n} \sum \log \frac{q(X_i)}{p(X_i)}$$

$$= -E(\log \frac{q(X)}{p(X)}) \text{ w.p. 1}$$

$$= - \sum p(x) \log \frac{q(x)}{p(x)}$$

$$= \sum p(x) \log \frac{p(x)}{q(x)}$$

$$= D(p||q).$$
10. **drawing a codebook** Let $X_i$ be a r.v. i.i.d distributed according to $P(x)$. We draw codebook of $2^nR$ codewords of $X^n$ independently using $P(x)$ and i.i.d.. We would like to answer the question: what is the probability that the first codeword would be identical to another codeword in the codebook as $n$ goes to infinity.

- Let $x^n$ be a sequence in the typical set $A^n(X)$. What is the asymptotic probability (you may provide an upper and lower bound) as $n \to \infty$ that we draw a sequence $X^n$ i.i.d distributed according to $P(x)$ and we get $x^n$.

- Using your answer from the previous sub-question find an $\alpha$ such that if $R < \alpha$ the probability that the first codeword in the codebook appears twice or more in the codebook goes to zero as $n \to \infty$.

- Find an $\alpha$ such that if $R > \alpha$ the probability that the first codeword in the codebook appears twice or more in the codebook goes to 1 as $n \to \infty$.

**Solution: drawing a codebook.**

(a) By the definition of the typical set, the probability of every $x^n \in A^n(e)$ to be drawn is bounded by

$$2^{-n(H(X)+\epsilon)} \leq p(x^n) \leq 2^{-n(H(X)-\epsilon)}$$

notice that the question regards to the probability of a specific sequence $x^n \in A^n(e)$ and not the probability of the whole set (which is almost 1).

(b) We want to find $\bar{\alpha}$ such that if $R < \bar{\alpha}$ then $\Pr(\exists i \neq 1 : x^n(i) =$
Consider the following derivations:

\[ \Pr(\exists i \neq 1 : x^n(i) = x^n(1)) = \Pr(\bigcup_{i=2}^{2nR} \{x^n(i) = x^n(1)\}) \]

\[ \leq \sum_{i=2}^{2nR} \Pr(x^n(i) = x^n(1)) \]

\[ \leq 2^{-n\alpha} \prod_{i=2}^{2nR} \left(1 - \Pr(x^n(i) = x^n(1))\right) \]

where (a) follows from the union bound and (b) follows from section a. If we set \( \alpha = H(X) - \epsilon \) this probability goes to zero.

(c) We want to find \( \alpha \) such that if \( R > \alpha \) then \( P(A) = 1 - \Pr(\forall i \neq 1 : x^n(i) \neq x^n(1)) \) goes to 1. Consider the following derivations:

\[ \Pr(\forall i \neq 1 : x^n(i) \neq x^n(1)) = \prod_{i=2}^{2nR} \Pr(x^n(i) \neq x^n(1)) \]

\[ \leq \prod_{i=2}^{2nR} (1 - \Pr(x^n(i) = x^n(1))) \]

\[ \leq \prod_{i=2}^{2nR} (1 - 2^{-n\alpha}) \]

\[ \leq \left(1 - 2^{-n\alpha}\right)^{2nR} \]

\[ \leq e^{-2nR \cdot n\alpha} \]

\[ \leq e^{-2nR - nH(X) + \epsilon} \]

where (a) follows from section a, (b) follows by adding \( i = 1 \) to the product and since the drawing is i.i.d, and (c) follows since \((1 - x)^n \leq e^{-nx}\). If we set \( R > H(X) + \epsilon \) this probability goes to zero and thus \( P(A) \) goes to 1. Hence, \( \alpha = H(X) + \epsilon \).

11. **Saving the princess.** A princess was abducted and was put in one of \( K \) rooms. Each room is labeled by a number 1, 2, \ldots, \( K \). Each room
is of size $s_i$ where $i = 1, 2, \ldots, K$. The probability of the princess to be in room $i$ is $p_i$, and proportional to the size of the room $s_i$, namely, $p_i = \alpha s_i$ where $\alpha$ is a constant.

(a) Find $\alpha$.

(b) In order to save the princess you need to find in which room she is. You may ask the demon a yes/no question. Like is she in room number 1 or is she in room 2 or 5 or is she in a room of odd number, and so on. You will save the princess if and only if the expected number of questions is the minimum possible. What would be the questions you should ask the demon in order to save the princess?

Solution: Saving the princess.

(a) Since $p$ is a probability mass function, it must satisfy

$$1 = \sum_{i=1}^{K} p_i = \sum_{i=1}^{K} \alpha s_i = \alpha \sum_{i=1}^{K} s_i$$

Therefore, $\alpha = \frac{1}{\sum_{i=1}^{K} s_i}$.

(b) The idea is to use Huffman code with the probability that the princess should be in a specific room. Follow the solution of Bad wine question. Here, we will build questions like is the princess in room 1 or 2 instead of mixing wine, and we want to find the princess instead of bad wine bottle.

12. Lossless source coding with side information.

Consider the lossless source coding with side information that is available at the encoder and decoder, where the source $X$ and the side information $Y$ are i.i.d. $\sim P_{X,Y}(x,y)$.

Show that a code with rate $R < H(X|Y)$ can not be achievable, and interpret the result.

Hint: Let $T \triangleq f(X^n, Y^n)$. Consider

$$nR \geq H(T) \geq H(T|Y^n),$$

(7)
and use similar steps, including Fano’s inequality, as we used in the
class to prove the converse where side information was not available.

**Solution** Sketch of the solution (please fill in the explanation for each
step):

\[
\begin{align*}
nR & \geq H(T) \\
& \geq H(T|Y^n), \\
& \geq I(X^n; T|Y^n) \\
& = H(X^n|Y^n) - H(X^n|T, Y^n) \\
& = nH(X|Y) - \epsilon_n,
\end{align*}
\]

where \( \epsilon_n \rightarrow 0 \).

13. **Challenge: Optimal code for an infinite alphabet** This question
is a challenge.

14. **Conditional Information Divergence**

(a) Let \( X, Z \) be random variables jointly distributed according to
\( P_{X,Z} \). We define the conditional informational divergence as fol-
follows:

\[
D(P_{X|Z}||Q_{X|Z}|P_Z) = \sum_{(x,z)\in X\times Z} P_{X,Z}(x, z) \log \left( \frac{P_{X|Z}(x|z)}{Q_{X|Z}(x|z)} \right).
\]

With respect to this definition, prove for each relation if it is true
or false:

For any pair of random variables \( A, B \) that are jointly distributed
according to \( P_{A,B} \),
\( D(P_{A,B} \mid Q_{A,B}) = D(P_A \mid Q_A) + D(P_{B|A} \mid Q_{B,A} \mid P_A). \)

ii. \( D(P_{A,B} \mid P_A P_B) = D(P_{B|A} \mid P_B \mid P_A). \)

iii. \( I(A; B) = D(P_{B|A} \mid P_B \mid P_A). \)

iv. \( D(P_{A|B} \mid Q_{A|B} \mid P_B) = \sum_{b \in B} P_B(b) D(P_{A|B=b} \mid Q_{A|B=b}). \)

(b) Consider the setting in Fig. 3.

Figure 3: Source coding with side information.

We would like to compress the source sequence \( X^n \) losslessly using a prefix code with side information \( Z^n \) which is available to the encoder and the decoder. The sources \( (X^n, Z^n) \) are distributed i.i.d. according to \( P_{X,Z} \) and that all the distribution and conditional distributions are dyadic (i.e., \( P_X \) is dyadic if \( P_X(x) = 2^{-i} \), for some \( i \), for all \( x \in X \)). We denote the average number of bits per symbol needed to compress the source \( X^n \) as \( L \).

i. What is the minimal \( L \)?

ii. Although the distribution of \( (X^n, Z^n) \) is \( P_{X,Z} \), the distribution that is used design the optimal prefix code is \( Q_{X|Z} P_Z \). What is the actual \( L \) (average bits per symbol) of this code?

iii. Now, the distribution that is used to design the prefix code is \( Q_{X,Z} \). What is the actual \( L \) now?

Solutions See question 4 in Final Exam 2014 Moed A (with solutions).
15. **True or False of a constrained inequality:**

Given are three discrete random variables $X, Y, Z$ that satisfy $H(Y|X, Z) = 0$.

(a) Copy the next relation and write **true** or **false** (If true, prove the statement, and if not provide a counterexample).

$$I(X; Y) \geq H(Y) - H(Z)$$

(b) What are the conditions for which the equality $I(X; Y) = H(Y) - H(Z)$ holds.

(c) Assume that the conditions for $I(X; Y) = H(Y) - H(Z)$ are satisfied. Is it true that there exists a function such that $Z = g(Y)$?

**Solution**

(a) First, note that $I(X; Y) = H(Y) - H(Y|X)$. Consider

$$0 = H(Y|X, Z)$$
$$= H(X, Y, Z) - H(X, Z)$$
$$= H(X) + H(Y|X) + H(Z|X, Y) - H(X) - H(Z|X)$$
$$= H(Y|X) - I(Z; Y|X)$$

Therefore, we can conclude that

$$H(Y|X) = I(Z; Y|X) \leq H(Z|X) \leq H(Z). \quad (8)$$

It follows that

$$H(Y) - H(Y|X) \geq H(Y) - H(Z).$$

(b) To satisfy equality, we must satisfy equalities in (8). First equality is obtained iff $H(Z|X, Y) = 0$, implies that $Z$ is a function of $X$ and $Y$. The second if $H(Z|X) = H(Z)$, which implies that $Z$ is a function of $X$.

(c) False. Consider the following counter example. Let $X$ and $Z$ be independent, each distributed according to Bernoulli($\frac{1}{2}$) and $Y = X \oplus Z$. Then, $Y = f(X, Z)$, $H(Z|X) = H(Z)$ and $Z = X \oplus Y$ so $H(Z|X, Y) = 0$. However, $Z$ is not a function of $Y$. 

18
16. **True or False**: Copy each relation and write true or false.

(a) Let $X \rightarrow Y \rightarrow Z \rightarrow W$ be a Markov chain, then the following holds:

$$I(X; W) \leq I(Y; Z).$$

(b) For two probability distributions, $p_{XY}$ and $q_{XY}$, that are defined on $\mathcal{X} \times \mathcal{Y}$, the following holds:

$$D(p_{XY}||q_{XY}) \geq D(p_X||q_X).$$

(c) If $X$ and $Y$ are dependent and also $Y$ and $Z$ are dependent, then $X$ and $Z$ are dependent.

**Solution**

(a) True. By data processing inequality, if $A \rightarrow B \rightarrow C$ form a Markov chain, then $I(A; C) \leq I(A; B)$. Here, we have $X \rightarrow Y \rightarrow Z$ and $X \rightarrow Z \rightarrow W$ as Markov chains.

$$I(X; W) \leq I(X; Z) \leq I(Y; Z).$$

(b) True. Consider the definition of a conditional divergence,

$$D(P_{X|Z}||Q_{X|Z}|P_Z) = \sum_{(x,z) \in \mathcal{X} \times \mathcal{Z}} P_{X,Z}(x, z) \log \left( \frac{P_{X|Z}(x|z)}{Q_{X|Z}(x|z)} \right).$$

From previous question on the conditional divergence, we learned that

$$D(P_{X,Y}||Q_{X,Y}) = D(P_X||Q_Y) + D(P_{Y|X}||Q_{Y|X}|P_X),$$

where

$$D(P_{Y|X}||Q_{Y|X}|P_X) = \sum_{x \in \mathcal{X}} P_x(x) D(P_{Y|X=x}||Q_{Y|X=x}),$$

which is non-negative. We conclude that

$$D(p_{XY}||q_{XY}) \geq D(p_X||q_X).$$

(c) False. Here is a counterexample. Let $X$ and $W$ be two independent random variables. Let $Y = X + W$ and $Z = W$. Then, $Y$ and $X$ are dependent, $Z$ and $Y$ are dependent, and $Z$ is independent of $X$. 

19